An Introduction to PAC-Bayesian Analysis

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Generalisation is the ability to 'perform' well on unseen data.

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- PAC: probably approximately correct [59]
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- PAC: probably approximately correct [59] Use a 'confidence parameter' δ : \mathbb{P}^m [large error] $\leq \delta$ δ is the probability of being misled by the training set
- Hence high confidence: \mathbb{P}^m [approximately correct] $\ge 1 \delta$

Error distribution picture



Learning algorithm $A : \mathcal{Z}^m \to \mathcal{H}$

• $\mathcal{Z} = \mathcal{X} \times \mathcal{Y}$ \mathcal{X} = set of inputs \mathcal{Y} = set of outputs (e.g. labels)

H = hypothesis class
 = set of predictors
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b these can be relaxed (mostly beyond the scope of this tutorial)

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Actually these two goals interact with each other!

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Theoretical risk: (out-of-sample)

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Examples:

- $\ell(h(X), Y) = \mathbf{1}[h(X) \neq Y]$: 0-1 loss (classification)
- $\ell(h(X), Y) = (Y h(X))^2$: square loss (regression)
- $\ell(h(X), Y) = (1 Yh(X))_+$: hinge loss
- $\ell(h(X), Y) = -\log(h(X))$: log loss (density estimation)

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Flavours:

- distribution-free
- algorithm-free

- distribution-dependent
- algorithm-dependent

Before PAC-Bayes
■ Single hypothesis *h* (building block):

with probability $\ge 1 - \delta$, $R_{\text{out}}(h) \le R_{\text{in}}(h) + \sqrt{\frac{1}{2m} \log(\frac{1}{\delta})}$.

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 \longrightarrow Extension: PAC-Bayes allows to consider *distributions* over hypotheses.

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The risk measures $R_{in}(h)$ and $R_{out}(h)$ are extended by averaging: $R_{in}(Q) \equiv \int_{\mathcal{H}} R_{in}(h) \, dQ(h) \qquad R_{out}(Q) \equiv \int_{\mathcal{H}} R_{out}(h) \, dQ(h)$

 $\operatorname{KL}(Q||P) = \underset{h \sim Q}{\mathsf{E}} \ln \frac{Q(h)}{P(h)}$ is the Kullback-Leibler divergence.

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Prior	Bayesian inference Unique Statistical modelling (likelihood)	Posterior
Any distribution	PAC-Bayes Model-free	Any distribution
on data	Inspired by the Bayesian update principle - Only depends on loss	on data

"Prior": exploration mechanism of ${\mathcal H}$ "Posterior" is the twisted prior after confronting with data

Prior

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Data distribution

- PAC-Bayes: bounds hold for any distribution
- · Bayes: randomness lies in the noise model generating the output

A General PAC-Bayesian Theorem

 Δ -function: "distance" between $R_{in}(Q)$ and $R_{out}(Q)$

Convex function $\Delta : [0, 1] \times [0, 1] \rightarrow \mathbb{R}$.

General theorem

(Bégin et al. [7, 8], Germain [21])

For any distribution D on $\mathfrak{X} \times \mathfrak{Y}$, for any set \mathfrak{H} of voters, for any distribution P on \mathfrak{H} , for any $\delta \in (0, 1]$, and for any Δ -function, we have, with probability at least $1-\delta$ over the choice of $S \sim D^m$,

$$\forall Q \text{ on } \mathcal{H}: \quad \Delta\Big(R_{\mathrm{in}}(Q), R_{\mathrm{out}}(Q)\Big) \leqslant \frac{1}{m} \Big[\mathrm{KL}(Q \| P) + \ln \frac{\mathtt{J}_{\Delta}(m)}{\delta}\Big],$$

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where

$$\mathcal{J}_{\Delta}(m) = \sup_{r \in [0,1]} \left[\sum_{k=0}^{m} \underbrace{\binom{m}{k} r^{k} (1-r)^{m-k}}_{\operatorname{Bin}(k;m,r)} e^{m\Delta(\frac{k}{m},r)} \right]$$

Proof of the general theorem

General theorem

$$\Pr_{\mathcal{S}\sim D^m}\left(\forall \, \mathcal{Q} \text{ on } \mathcal{H}: \, \Delta\left(R_{\mathrm{in}}(\mathcal{Q}), R_{\mathrm{out}}(\mathcal{Q})\right) \leq \frac{1}{m}\left[\mathrm{KL}(\mathcal{Q}\|\mathcal{P}) + \ln \frac{\mathtt{J}_{\Delta}(m)}{\delta}\right]\right) \, \geq \, 1-\delta \, .$$

Proof ideas.

Change of Measure Inequality

For any P and Q on $\mathcal H,$ and for any measurable function $\varphi:\mathcal H\to\mathbb R,$ we have

$$-\ln\left(\mathop{\mathsf{E}}_{h\sim P} e^{\Phi(h)}\right) = -\ln\mathop{\mathsf{E}}_{h\sim Q}\left(\frac{P(h)}{Q(h)}e^{\Phi(h)}\right)$$
$$\leqslant \mathop{\mathsf{E}}_{h\sim Q}\ln\left(\frac{Q(h)}{P(h)}\right) - \mathop{\mathsf{E}}_{h\sim Q}\Phi(h)$$
$$= \operatorname{KL}(Q||P) - \mathop{\mathsf{E}}_{h\sim Q}\Phi(h).$$

Proof of the general theorem

General theorem

$$\Pr_{S \sim D^m} \left(\forall Q \text{ on } \mathcal{H} : \Delta\left(R_{\text{in}}(Q), R_{\text{out}}(Q)\right) \leq \frac{1}{m} \left[\text{KL}(Q \| P) + \ln \frac{J_{\Delta}(m)}{\delta} \right] \right) \geq 1 - \delta.$$

Proof ideas.

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$$\leqslant \underset{h\sim Q}{\mathsf{E}}\ln\left(\frac{Q(h)}{P(h)}\right) - \underset{h\sim Q}{\mathsf{E}}\phi(h)$$
$$= \operatorname{KL}(Q||P) - \underset{h\sim Q}{\mathsf{E}}\phi(h).$$

Markov's inequality

for a random variable X satisfying $X \ge 0$ $\Pr(X \ge a) \le \frac{\mathbb{E}X}{a} \iff \Pr(X \le \frac{\mathbb{E}X}{\delta}) \ge 1 - \delta$.

Proof of the general theorem

Probability of observing k misclassifications among m examples Given a voter h, consider a **binomial variable** of m trials with **success** $R_{out}(h)$:

$$\Pr_{S\sim D^m}\left(R_{\rm in}(h) = \frac{k}{m}\right) = \binom{m}{k} \left(R_{\rm out}(h)\right)^k \left(1 - R_{\rm out}(h)\right)^{m-k} = \operatorname{Bin}\left(k; m, R_{\rm out}(h)\right)$$

$$m \cdot \Delta \Big(\underset{h \sim Q}{\mathsf{E}} R_{\mathrm{in}}(h), \underset{h \sim Q}{\mathsf{E}} R_{\mathrm{out}}(h) \Big)$$

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Jensen's Inequality

 \leq

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Jensen's Inequality
$$\leq \qquad \underset{h \sim Q}{\mathsf{E}} m \cdot \Delta \left(R_{\mathrm{in}}(h), R_{\mathrm{out}}(h) \right)$$
Change of measure
$$\leq \qquad \mathrm{KL}(Q \| P) + \ln \underset{h \sim P}{\mathsf{E}} e^{m\Delta \left(R_{\mathrm{in}}(h), R_{\mathrm{out}}(h) \right)}$$

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General theorem

$$\Pr_{\mathcal{S}\sim D^m}\left(\forall Q \text{ on } \mathcal{H}: \Delta\left(R_{\text{in}}(Q), R_{\text{out}}(Q)\right) \leq \frac{1}{m}\left[\text{KL}(Q||P) + \ln \frac{J_{\Delta}(m)}{\delta}\right]\right) \geq 1-\delta.$$

Corollary

[...] with probability at least $1-\delta$ over the choice of $S \sim D^m$, for all Q on \mathcal{H} :

(a) $\operatorname{kl}\left(R_{\operatorname{in}}(Q), R_{\operatorname{out}}(Q)\right) \leq \frac{1}{m}\left[\operatorname{KL}(Q||P) + \ln \frac{2\sqrt{m}}{\delta}\right]$, Langford and Seeger [31]

$$\frac{\mathrm{kl}(q,\rho)}{=} \quad q \ln \frac{q}{\rho} + (1-q) \ln \frac{1-q}{1-\rho}$$

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 $[C] R_{\text{out}}(Q) \leq \frac{1}{1 - e^{-c}} \left(c \cdot R_{\text{in}}(Q) + \frac{1}{m} \left[\text{KL}(Q \| P) + \ln \frac{1}{\delta} \right] \right),$

Catoni [11]

$$\begin{split} & \mathrm{kl}(q,p) & \stackrel{\mathrm{def}}{=} & q \ln \frac{q}{p} + (1-q) \ln \frac{1-q}{1-p} \geqslant 2(q-p)^2 \,, \\ & \Delta_c(q,p) & \stackrel{\mathrm{def}}{=} & -\ln[1-(1-e^{-c})\cdot p] - c \cdot q \,, \end{split}$$
General theorem

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Corollary

 $\begin{array}{ll} [\dots] \text{ with probability at least } 1-\delta \text{ over the choice of } S \sim D^m, \text{ for all } Q \text{ on } \mathcal{H}: \\ \hline \textbf{a} & \mathrm{kl}\Big(R_{\mathrm{in}}(Q), R_{\mathrm{out}}(Q)\Big) \leq \frac{1}{m} \left[\mathrm{KL}(Q||P) + \ln \frac{2\sqrt{m}}{\delta}\right], & \text{Langford and Seeger [31]} \\ \hline \textbf{b} & R_{\mathrm{out}}(Q) \leq R_{\mathrm{in}}(Q)) + \sqrt{\frac{1}{2m} \left[\mathrm{KL}(Q||P) + \ln \frac{2\sqrt{m}}{\delta}\right]}, & \text{McAllester [40, 43]} \\ \hline \textbf{c} & R_{\mathrm{out}}(Q) \leq \frac{1}{1-e^{-c}} \left(c \cdot R_{\mathrm{in}}(Q) + \frac{1}{m} \left[\mathrm{KL}(Q||P) + \ln \frac{1}{\delta}\right]\right), & \text{Catoni [11]} \\ \hline \textbf{c} & R_{\mathrm{out}}(Q) \leq R_{\mathrm{in}}(Q) + \frac{1}{\lambda} \left[\mathrm{KL}(Q||P) + \ln \frac{1}{\delta} + f(\lambda, m)\right]. & \text{Alquier et al. [4]} \end{array}$

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$$\begin{split} \mathbf{E}_{S \sim D^{m} \ h \sim P} \mathbf{E}_{m} \mathbf{E}_{h \sim P} e^{m\Delta(R_{S}(h), R(h))} &= \mathbf{E}_{h \sim P} \mathbf{E}_{S \sim D^{m}} \left(\frac{R_{S}(h)}{R(h)}\right)^{mR_{S}(h)} \left(\frac{1-R_{S}(h)}{1-R(h)}\right)^{m(1-R_{S}(h))} \\ &= \mathbf{E}_{h \sim P} \sum_{k=0}^{m} \Pr_{S \sim D^{m}} \left(R_{S}(h) = \frac{k}{m}\right) \left(\frac{k}{R(h)}\right)^{k} \left(\frac{1-\frac{k}{m}}{1-R(h)}\right)^{m-k} \\ &= \sum_{k=0}^{m} {m \choose k} (k/m)^{k} (1-k/m)^{m-k}, \end{split}$$
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$$&\leqslant 2\sqrt{m}. \end{split}$$

■ Note that, in Line (1) of the proof, $\Pr_{S \sim D^m} (R_S(h) = \frac{k}{m})$ is replaced by the probability mass function of the binomial.

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- So this result is no longer valid in the non iid case, even if General Theorem is.

Linear classifiers

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- The specification of the centre for the posterior Q(w, μ) will be by a unit vector mw and a scale factor μ.









Linear classifiers performance may be bounded by

$$\mathsf{KL}(\hat{Q}_{\mathcal{S}}(\mathbf{w},\mu) \| \mathbf{Q}_{\mathcal{D}}(\mathbf{w},\mu)) \leqslant \frac{\mathsf{KL}(P \| Q(\mathbf{w},\mu)) + \ln \frac{m+1}{\delta}}{m}$$

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■ SVM is deterministic classifier that exactly corresponds to sgn (E_{c~Q(mw,µ)}[c(x)]) as centre of the Gaussian gives the same classification as halfspace with more weight.

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- SVM is deterministic classifier that exactly corresponds to sgn (E_{c~Q(mw,µ)}[c(x)]) as centre of the Gaussian gives the same classification as halfspace with more weight.
- Hence its error bounded by $2Q_{D}(\mathbf{m}w, \mu)$, since as observed above if **x** misclassified at least half of $c \sim Q$ err.

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Â_S(**w**, μ) stochastic measure of the training error
Â_S(**w**, μ) = E_m[F̃(μγ(**x**, y))]

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 $\hat{Q}_{S}(\mathbf{w}, \mu) \text{ stochastic measure of the training error}$ $\hat{Q}_{S}(\mathbf{w}, \mu) = \mathbb{E}_{m}[\tilde{F}(\mu\gamma(\mathbf{x}, y))]$ $\mathbf{v}(\mathbf{x}, y) = (y\mathbf{w}^{T} \Phi(\mathbf{x}))/(\|\Phi(\mathbf{x})\|\|\mathbf{w}\|)$

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- The bound holds with probability 1 δ over the random i.i.d. selection of the training data.

Form of the SVM bound

Note that bound holds for all posterior distributions so that we can choose μ to optimise the bound

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- Note that bound holds for all posterior distributions so that we can choose μ to optimise the bound
- If we define the inverse of the KL by

 $\mathrm{KL}^{-1}(q, A) = \max\{p : \mathrm{KL}(q \| p) \leqslant A\}$

then have with probability at least $1-\delta$

$$\Pr\left(\langle \mathbf{w}, \boldsymbol{\varphi}(\mathbf{x}) \rangle \neq y\right) \leqslant 2\min_{\mu} \mathrm{KL}^{-1}\left(\mathbb{E}_m[\tilde{F}(\mu\gamma(\mathbf{x}, y))], \frac{\mu^2/2 + \ln\frac{m+1}{\delta}}{m}\right)$$

Gives SVM Optimisation

Primal form:

$$\min_{\mathbf{w},\xi_i} \begin{bmatrix} \frac{1}{2} \|\mathbf{w}\|^2 + C \sum_{i=1}^m \xi_i \end{bmatrix}$$

s.t. $y_i \mathbf{w}^T \phi(\mathbf{x}_i) \ge 1 - \xi_i$ $i = 1, \dots, m$
 $\xi_i \ge 0$ $i = 1, \dots, m$

$$\max_{\alpha} \left[\sum_{i=1}^{m} \alpha_{i} - \frac{1}{2} \sum_{i,j=1}^{m} \alpha_{i} \alpha_{j} y_{i} y_{j} \kappa(\mathbf{x}_{i}, \mathbf{x}_{j}) \right]$$

s.t. $0 \leq \alpha_{i} \leq C \quad i = 1, ..., m$

where $\kappa(\mathbf{x}_i, \mathbf{x}_j) = \langle \varphi(\mathbf{x}_i), \varphi(\mathbf{x}_j) \rangle$ and $\langle \mathbf{w}, \varphi(\mathbf{x}) \rangle = \sum_{i=1}^m \alpha_i y_i \kappa(\mathbf{x}_i, \mathbf{x}).$



Model Selection with the new bound: setup

 Comparison of 10-fold Xvalidation, PAC-Bayes Bound and the Prior PAC-Bayes Bound

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 - For PAC-Bayes Bound and Prior PAC-Bayes Bound select the pair that minimize the bound

Results

		Classifier					
		SVM				ηPrior SVM	
Problem		2FCV	10FCV	PAC	PrPAC	PrPAC	τ-PrPAC
digits	Bound	_	_	0.175	0.107	0.050	0.047
	TE	0.007	0.007	0.007	0.014	0.010	0.009
waveform	Bound	_	_	0.203	0.185	0.178	0.176
	TE	0.090	0.086	0.084	0.088	0.087	0.086
pima	Bound	_	_	0.424	0.420	0.428	0.416
	TE	0.244	0.245	0.229	0.229	0.233	0.233
ringnorm	Bound	_	_	0.203	0.110	0.053	0.050
	TE	0.016	0.016	0.018	0.018	0.016	0.016
spam	Bound	_	_	0.254	0.198	0.186	0.178
	TE	0.066	0.063	0.067	0.077	0.070	0.072
Average	TE	0.0846	0.0834	0.081	0.0852	0.0832	0.0832

Take home messages

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- Bounds are remarkably tight: for final column average factor between bound and TE is under 3.
- Model selection from the bounds is as good as 10FCV: in fact all but one of the PAC-Bayes model selections give better averages for TE.
- The better bounds do not appear to give better model selection best model selection is from the simplest bound.
 - A. Ambroladze, E. Parrado-Hernández, and J. Shawe-Taylor. Tighter PAC-Bayes bounds. In *Advances in Neural Information Processing Systems* 18, (2006) Pages 9-16.
 - P. Germain, A. Lacasse, F. Laviolette and M. Marchand. PAC-Bayesian learning of linear classifiers, in *Proceedings of the 26nd International Conference on Machine Learning* (ICML'09, Montréal, Canada.). ACM Press (2009), 382, Pages 453-460.

Deep Learning Results



Deep Learning Results



A flexible framework

A flexible framework

Since 1997, PAC-Bayes has been successfully used in many machine learning settings (this list is by no means exhaustive).

Statistical learning theory Audibert and Bousquet [6], Catoni [9, 10], Guedj [25], Guedj and Pujol [27], Maurer [39], McAllester [41, 42, 44, 45], Mhammedi et al. [46], Seeger [51, 52], Shawe-Taylor

and Williamson [56], Thiemann et al. [58]

- SVMs & linear classifiers Germain et al. [19], Langford and Shawe-Taylor [32], McAllester [44]
- Supervised learning algorithms reinterpreted as bound minimizers Ambroladze et al. [5], Germain et al. [22], Shawe-Taylor and Hardoon [57]
- High-dimensional regression Alquier and Biau [1], Alquier and Lounici [2], Guedj and Robbiano [24], Guedj and Alquier [26], Li et al. [35]
 Classification Catoni [9, 10], Lacasse et al. [30], Langford and Shawe-Taylor [32], Parrado-Hernández et al. [49]

A flexible framework

Transductive learning, domain adaptation Bégin et al. [7]. Derbeko et al. [12], Germain et al. [20], Nozawa et al. [48] Non-iid or heavy-tailed data Alquier and Guedj [3], Holland [29], Lever et al. [34], Seldin et al. [54, 55] Density estimation Higgs and Shawe-Taylor [28], Seldin and Tishby [53] Reinforcement learning Fard and Pineau [16], Fard et al. [17], Ghavamzadeh et al. [23], Seldin et al. [54, 55] Sequential learning Gerchinovitz [18], Li et al. [36] Algorithmic stability, differential privacy Dziugaite and Roy [13, 14], London [37], London et al. [38], Rivasplata et al. [50] Deep neural networks Dziugaite and Roy [15], Letarte et al. [33], Neyshabur

et al. [47], Zhou et al. [60]

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